Goodness-of-fit tests based on Wasserstein distance to detect changes on local protein conformations

Javier González-Delgado

Institut de Mathématiques de Toulouse, LAAS-CNRS

Journée toulousaine "Statistique pour la Biologie"

November 15, 2022

Joint work with

Pierre Neuvial¹ Juan Cortés² Pau Bernadó³ Alberto González-Sanz^{1,4}

1. Institut de Mathématiques de Toulouse, Université de Toulouse, CNRS, Toulouse, France. 2. LAAS-CNRS, Université de Toulouse, CNRS, Toulouse, France. 3. Centre de Biologie Structurale, Université de Montpellier, INSERM, CNRS, France. 4. ImUva, Universidad de Valladolid.

Primary structure

Primary structure

Secondary structure (conformation)

Primary structure

Secondary structure (conformation)

Sequence $\stackrel{?}{\Leftrightarrow}$ 3D structure \Leftrightarrow Function

Javier González-Delgado **Masserstein tests for IDP** 15/11/2022 3/20

At each amino-acid, local conformation is given by the torsion dihedral angles $¹$ </sup>

- $\cdot \omega_i \in \{0, \pi\}$ is fixed,
- \cdot $(\varphi_i, \psi_i) \in [-\pi, \pi] \times [-\pi, \pi]$ determine local conformation.

1. G.N. RAMACHANDRAN et al. "Stereochemistry of polypeptide chain configurations". In : Journal of Molecular Biology 7.1 (1963), p. 95-99

(φ, ψ) dihedral angles

- · are physically restricted for globular (structured) proteins,
- · are random for Intrinsically Disordered Proteins (IDP).

Local conformation is given by a probability measure supported on \mathbb{T}^2 .

Detecting changes on local protein conformations

Detecting changes on local protein conformations \mathcal{D}

Detecting changes on probability measures supported on \mathbb{T}^2

Detecting changes on local protein conformations \hat{U} Detecting changes on probability measures supported on \mathbb{T}^2 \hat{U} Goodness-of-fit testing for measures supported on \mathbb{T}^2

Detecting changes on local protein conformations $\hat{\mathbb{I}}$ Detecting changes on probability measures supported on \mathbb{T}^2 \mathcal{D} Goodness-of-fit testing for measures supported on \mathbb{T}^2

First question : choice of a metric between distributions (statistic).

Optimal Transport Theory Wasserstein distance

- · C. Villani. Topics in Optimal Transportation. Providence, Rhode Island : American mathematical society, 2003
- · C. Villani. Optimal Transport : Old and New. Springer-Verlag Berlin Heidelberg, 2008

Optimal Transport Theory Wasserstein distance

- · C. Villani. Topics in Optimal Transportation. Providence, Rhode Island : American mathematical society, 2003
- · C. Villani. Optimal Transport : Old and New. Springer-Verlag Berlin Heidelberg, 2008

p-Wasserstein distance between two arbitrary measures

$$
\mathcal{W}_p^p(\mu,\nu) = \min_{\pi \in \mathcal{U}(\mu,\nu)} \int_{\mathcal{X} \times \mathcal{Y}} c(x,y)^p \mathrm{d} \pi(x,y) = \min_{(X,Y)} \left\{ \mathbb{E}_{(X,Y)}(c(X,Y)^p) : X \sim \mu \ Y \sim \nu \right\}.
$$

Goodness-of-fit test

Let $P,Q\in\mathcal{P}(\mathbb{R}^2/\mathbb{Z}^2)$ and P_n,Q_m the corresponding empirical probability measures. The goal is to assess the null hypothesis :

$$
H_0: P = Q \tag{1}
$$

by considering the p-value

$$
\mathbb{P}_{H_0}(\mathcal{W}_p^p(P_n,Q_m)\geq w_{nm}),\qquad \qquad (2)
$$

where w_{nm} is the statistic realization.

Goodness-of-fit test

Let $P,Q\in\mathcal{P}(\mathbb{R}^2/\mathbb{Z}^2)$ and P_n,Q_m the corresponding empirical probability measures. The goal is to assess the null hypothesis :

$$
H_0: P = Q \tag{1}
$$

by considering the p-value

$$
\mathbb{P}_{H_0}(\mathcal{W}_p^p(P_n,Q_m)\geq w_{nm}),\qquad \qquad (2)
$$

where w_{nm} is the statistic realization.

Problem

We need to know the null distribution of $\mathcal{W}^p_p(P_n,Q_m).$

Goodness-of-fit test

Let $P,Q\in\mathcal{P}(\mathbb{R}^2/\mathbb{Z}^2)$ and P_n,Q_m the corresponding empirical probability measures. The goal is to assess the null hypothesis :

$$
H_0: P = Q \tag{1}
$$

by considering the p-value

$$
\mathbb{P}_{H_0}(\mathcal{W}_p^p(P_n,Q_m)\geq w_{nm}),\qquad \qquad (2)
$$

where w_{nm} is the statistic realization.

Problem

We need to know the null distribution of $\mathcal{W}^p_p(P_n,Q_m).$

Exact tests are unfeasible when dimension is higher than one

Distribution of $W_p^p(P_n, Q_m)$ when $P = Q$ is unknown.

Testing the equality of N_g projections to closed geodesics²

^{2.} J. GONZÁLEZ-DELGADO et al. Two-sample goodness-of-fit tests on the flat torus based on Wasserstein distance and their relevance to structural biology. arXiv :2108.00165. 2021

Testing the equality of N_g projections to closed geodesics²

Strategy

For each $i \in \{1, \ldots, N_{\epsilon}\}\;$:

- 1 Project both samples to the i-th closed geodesic,
- 2 Test the equality of both projected samples : get the i -th p -value.

Finally : aggregate the N_g p-values.

^{2.} J. GONZÁLEZ-DELGADO et al. Two-sample goodness-of-fit tests on the flat torus based on Wasserstein distance and their relevance to structural biology. arXiv :2108.00165. 2021

Testing the equality of $N_{\rm g}$ projections to closed geodesics²

Strategy

For each $i \in \{1, \ldots, N_{\epsilon}\}\;$:

- 1 Project both samples to the i-th closed geodesic,
- 2 Test the equality of both projected samples : get the i -th p -value.

Finally : aggregate the N_g p-values.

Relevant points

- How to choose geodesics : deterministic/random (uniform ?) sampling,
- **•** How to project samples to a given geodesic.

^{2.} J. GONZÁLEZ-DELGADO et al. Two-sample goodness-of-fit tests on the flat torus based on Wasserstein distance and their relevance to structural biology. arXiv :2108.00165. 2021

Geodesics on \mathbb{T}^2

Geodesics 3 on \mathbb{T}^2 are the images by the canonical projection of straight lines on $\mathbb{R}^2.$ Lines with irrational slope map to geodesics which are dense on $\mathbb{T}^2,$ and only lines with rational slope map to closed geodesics on the torus, which are closed spirals isomorphic to \mathbb{R}/\mathbb{Z} .

^{3.} William M. BOOTHBY. An Introduction to Differentiable Manifolds and Riemannian Geometry. Pure and Applied Mathematics. Academic Press, London, 1975

Using Wasserstein distance as test statistic

Let $P^c, Q^c \in \mathcal{P}(\R/\Z)$ and P^c_n, Q^c_m be their corresponding empirical probability measures. We aim to test

 $H_0: P^c = Q^c$ against $H_1: P^c \neq Q^c$.

^{4.} Julie DELON et al. "Fast transport optimization for Monge cost on the circle". In: SIAM Journal on Applied Mathematics 70.7/8 (2010), p. 2239-2258

Using Wasserstein distance as test statistic

Let $P^c, Q^c \in \mathcal{P}(\R/\Z)$ and P^c_n, Q^c_m be their corresponding empirical probability measures. We aim to test

$$
H_0: P^c = Q^c \qquad \text{against} \qquad H_1: P^c \neq Q^c.
$$

If we denote by F, G the cumulative distribution functions of P^c , Q^c respectively, defined as $F(t) = P^c([0,t])$. Then, we have⁴

$$
W_2^2(P^c, Q^c) = \inf_{\alpha \in \mathbb{R}} \int_0^1 (F^{-1}(t) - (G - \alpha)^{-1}(t))^2 dt,
$$
 (3)

where the pseudo-inverse is defined as $\mathit{F}^{-1}(s) = \inf\{t\,:\, \mathit{F}(t) > s\}.$

^{4.} Julie DELON et al. "Fast transport optimization for Monge cost on the circle". In: SIAM Journal on Applied Mathematics 70.7/8 (2010), p. 2239-2258

Using Wasserstein distance as test statistic

Let $P^c, Q^c \in \mathcal{P}(\R/\Z)$ and P^c_n, Q^c_m be their corresponding empirical probability measures. We aim to test

 $H_0: P^c = Q^c$ against $H_1: P^c \neq Q^c$.

If we denote by F, G the cumulative distribution functions of P^c , Q^c respectively, defined as $F(t) = P^c([0,t])$. Then, we have⁴

$$
W_2^2(P^c, Q^c) = \inf_{\alpha \in \mathbb{R}} \int_0^1 (F^{-1}(t) - (G - \alpha)^{-1}(t))^2 dt,
$$
 (3)

where the pseudo-inverse is defined as $\mathit{F}^{-1}(s) = \inf\{t\,:\, \mathit{F}(t) > s\}.$

Consequence

The Optimal Transport problem on \mathbb{R}/\mathbb{Z} reduces to the same problem on [0, 1) if both measures are relocated on the real line choosing as origin the minimizing element α .

^{4.} Julie DELON et al. "Fast transport optimization for Monge cost on the circle". In: SIAM Journal on Applied Mathematics 70.7/8 (2010), p. 2239-2258

Using Wasserstein distance as test statistic

Let $P^c, Q^c \in \mathcal{P}(\R/\Z)$ and P^c_n, Q^c_m be their corresponding empirical probability measures. We aim to test

 $H_0: P^c = Q^c$ against $H_1: P^c \neq Q^c$.

If we denote by F, G the cumulative distribution functions of P^c , Q^c respectively, defined as $F(t) = P^c([0,t])$. Then, we have⁴

$$
W_2^2(P^c, Q^c) = \inf_{\alpha \in \mathbb{R}} \int_0^1 (F^{-1}(t) - (G - \alpha)^{-1}(t))^2 dt,
$$
 (3)

where the pseudo-inverse is defined as $\mathit{F}^{-1}(s) = \inf\{t\,:\, \mathit{F}(t) > s\}.$

Consequence

The Optimal Transport problem on \mathbb{R}/\mathbb{Z} reduces to the same problem on $[0,1)$ if both measures are relocated on the real line choosing as origin the minimizing element α .

But... $W_2^2(P^c, Q^c)$ is not distribution-free under H_0 .

^{4.} Julie DELON et al. "Fast transport optimization for Monge cost on the circle". In: SIAM Journal on Applied Mathematics 70.7/8 (2010), p. 2239-2258

Finding a distribution-free statistic (I)

Adapting the idea from Ramdas et. al 5 to \mathbb{R}/\mathbb{Z}

Instead of comparing \digamma to G , compare $G(\digamma^{-1})$ to a uniform distribution.

^{5.} Aaditya RAMDAS et al. "On Wasserstein Two Sample Testing and Related Families of Nonparametric Tests". In: Entropy 19 (sept. 2015). poi : [10.3390/e19020047](https://doi.org/10.3390/e19020047)

Finding a distribution-free statistic (I)

Adapting the idea from Ramdas et. al 5 to \mathbb{R}/\mathbb{Z}

Instead of comparing \digamma to G , compare $G(\digamma^{-1})$ to a uniform distribution.

Two advantages :

- **Explicit form of the minimizer** α **.**
- \bullet Distribution-free statistic under H_0 .

^{5.} Aaditya RAMDAS et al. "On Wasserstein Two Sample Testing and Related Families of Nonparametric Tests". In: Entropy 19 (sept. 2015). poi : [10.3390/e19020047](https://doi.org/10.3390/e19020047)

Finding a distribution-free statistic (I)

Adapting the idea from Ramdas et. al 5 to \mathbb{R}/\mathbb{Z}

Instead of comparing \digamma to G , compare $G(\digamma^{-1})$ to a uniform distribution.

Two advantages :

- **Explicit form of the minimizer** α **.**
- \bullet Distribution-free statistic under H_0 .

Lemma (Explicit form of α)

Let $P^c \in \mathcal{P}(\mathbb{R}/\mathbb{Z})$, and F be its cumulative distribution function. Let U be the uniform distribution on R/Z. Then,

$$
W_2^2(P^c, U) = \int_0^1 (F^{-1}(t) - t - \alpha_0(F))^2 dt,
$$

where the optimal origin is given by $\alpha_0(\mathcal{F}) = \int_0^1 (\mathcal{F}^{-1}(t)-t) \, dt.$

^{5.} Aaditya RAMDAS et al. "On Wasserstein Two Sample Testing and Related Families of Nonparametric Tests". In: Entropy 19 (sept. 2015). DOI: [10.3390/e19020047](https://doi.org/10.3390/e19020047)

Two-sample goodness-of-fit test for measures on \mathbb{R}/\mathbb{Z} Finding a distribution-free statistic (II)

We consider the statistic

$$
T_{nm}^c = \frac{nm}{n+m}W_2^2(G_m \# P_n^c, U) = \frac{nm}{n+m} \int_0^1 (G_m(F_n^{-1}(t)) - t - \alpha_0(F_n^{-1}(G_m)))^2 dt,
$$

where $G_m \# P_n^c$ is the *push-forward* measure of P_n^c through G_m , and whose CDF is $G_m(F_n^{-1})$.

Two-sample goodness-of-fit test for measures on \mathbb{R}/\mathbb{Z} Finding a distribution-free statistic (II)

We consider the statistic

$$
T_{nm}^c = \frac{nm}{n+m}W_2^2(G_m \# P_n^c, U) = \frac{nm}{n+m} \int_0^1 (G_m(F_n^{-1}(t)) - t - \alpha_0(F_n^{-1}(G_m)))^2 dt,
$$

where $G_m \# P_n^c$ is the *push-forward* measure of P_n^c through G_m , and whose CDF is $G_m(F_n^{-1})$.

Proposition (Distribution free under H_0)

Let $P^c, Q^c \in \mathcal{P}(\mathbb{R}/\mathbb{Z})$, P_n^c, Q_m^c be their corresponding empirical probability measures, and F_n , G_m be their empirical cumulative distribution functions. If $\frac{n}{m} \to \lambda$ when $n, m \to \infty$ for some $\lambda \in [0, \infty)$ then, under $P^c = Q^c$, it holds that

$$
T_{nm}^c=\frac{nm}{n+m}\mathcal{W}_2^2(G_m\#P_n^c,U)\underset{n,m}{\xrightarrow{w}}\int_0^1\mathbb{B}(t)^2 dt-\left(\int_0^1\mathbb{B}(t) dt\right)^2,
$$

where $\mathbb B$ is a standard Brownian bridge, and the weak convergence is understood as convergence of probability measures on the space of right-continuous functions with left limits.

Two-sample test for measures on \mathbb{R}/\mathbb{Z}

$$
\pi_{nm}^c = \left\{ \begin{array}{ll} 1 & \textrm{if} & T_{nm}^c \ge c_{nm}^c(\alpha) \\ 0 & \textrm{otherwise} \end{array} \right.
$$

where the critical value $c_{nm}^{c}(\alpha)$ is given by

$$
c_{nm}^c(\alpha) = \inf\left\{t>0\,:\,F_{nm}^c(t)\geq 1-\alpha\right\}\,,
$$

with F_{nm}^c denoting the distribution function of T_{nm}^c under H_0 . Equivalently, a p -value for this test is $p_{nm}^c = 1 - F_{nm}^c(T_{nm}^c)$.

Two-sample test for measures on \mathbb{R}/\mathbb{Z}

$$
\pi_{nm}^c = \left\{ \begin{array}{ll} 1 & \textrm{if} & T_{nm}^c \ge c_{nm}^c(\alpha) \\ 0 & \textrm{otherwise} \end{array} \right.
$$

where the critical value $c_{nm}^{c}(\alpha)$ is given by

$$
c_{nm}^c(\alpha) = \inf\left\{t>0\,:\,F_{nm}^c(t)\geq 1-\alpha\right\}\,,
$$

with F_{nm}^c denoting the distribution function of T_{nm}^c under H_0 . Equivalently, a p -value for this test is $p_{nm}^c = 1 - F_{nm}^c(T_{nm}^c)$.

Proposition (Consistency)

Let $P^c, Q^c \in \mathcal{P}(\mathbb{R}/\mathbb{Z})$. If $P^c \neq Q^c$, it holds

$$
\lim_{n,m\to\infty} \mathbb{P}\left(\pi_{nm}^c=1\right)=1 \quad \text{for any } \alpha>0.
$$

Aggregating N_g *p*-values

Let p_i be the *i*-th p-value, for $i = 1, ..., N_g$. Then, under H_0 and assuming that the p-values are independent, the statistic

$$
T_{nm,N_g}^g=\min_{i=1}^{N_g}p_i
$$

follows a $\beta(1, N_g)$ distribution.

Aggregating N_g *p*-values

Let p_i be the *i*-th p-value, for $i = 1, \ldots, N_g$. Then, under H_0 and assuming that the p-values are independent, the statistic

$$
T_{nm,N_g}^g=\min_{i=1}^{N_g} p_i
$$

follows a $\beta(1, N_g)$ distribution. Consequently, the random variable

$$
\digamma_{\beta(1,N_g)}\left(\mathcal{T}^g_{nm,N_g}\right),
$$

where $F_{\beta(1,N_{\sigma})}$ denotes the $\beta(1,N_g)$ CDF, follows a uniform distribution on [0, 1].

Aggregating $N_{\rm g}$ p-values

Let p_i be the *i*-th p-value, for $i = 1, ..., N_g$. Then, under H_0 and assuming that the p-values are independent, the statistic

$$
T_{nm,N_g}^g=\min_{i=1}^{N_g} p_i
$$

follows a $\beta(1, N_{\epsilon})$ distribution. Consequently, the random variable

$$
\mathsf{F}_{\beta(1,\mathsf{N}_g)}\left(\mathsf{T}_{nm,\mathsf{N}_g}^g\right),\,
$$

where $F_{\beta(1,N_{\sigma})}$ denotes the $\beta(1,N_g)$ CDF, follows a uniform distribution on [0, 1].

Two-sample test for measures on \mathbb{T}^2

$$
\pi^g_{nm,N_g} = \begin{cases} 1 & \text{if} & F_{\beta(1,N_g)}\left(T^g_{nm,N_g}\right) \le \alpha \\ 0 & \text{otherwise} \end{cases} \qquad (N_g\text{-geod})
$$

Aggregating $N_{\rm g}$ p-values

Let p_i be the *i*-th p-value, for $i = 1, ..., N_g$. Then, under H_0 and assuming that the p-values are independent, the statistic

$$
T_{nm,N_g}^g=\min_{i=1}^{N_g} p_i
$$

follows a $\beta(1, N_g)$ distribution. Consequently, the random variable

$$
F_{\beta(1,N_g)}\left(T^g_{nm,N_g}\right),\,
$$

where $F_{\beta(1,N_{\sigma})}$ denotes the $\beta(1,N_g)$ CDF, follows a uniform distribution on [0, 1].

Two-sample test for measures on \mathbb{T}^2

$$
\pi^g_{nm,N_g} = \left\{ \begin{array}{ll} 1 & \textrm{if} & F_{\beta(1,N_g)} \left(T^g_{nm,N_g} \right) \leq \alpha \\ 0 & \textrm{otherwise} \end{array} \right. \qquad \qquad (N_g\text{-geod})
$$

Proposition (Consistency)

Let $P, Q \in \mathcal{P}(\mathbb{T}^2)$ and P_i^c (resp. Q_i^c), $i = 1, ..., N_g$, be the circular projected distributions of P (resp. Q) to N_g closed geodesics of \mathbb{T}^2 . If $P^c_i\neq Q^{\bar{c}}_i$ for at least one $i\in\{1,\ldots,N_g\},$ it holds

$$
\lim_{n,m\to\infty} \mathbb{P}\left(\pi^g_{nm,N_g}=1\right) \quad \text{for any } \alpha>0.
$$

Power analysis

Alternatives converging to the null (Uniform on \mathbb{T}^2 vs. bivariate von Mises converging to uniform)

Power analysis

Alternatives converging to the null (Uniform on \mathbb{T}^2 vs. bivariate von Mises converging to uniform)

Sample size (n = m) = W−geodesic (Ng = 2) − W−geodesic (Ng = 4) − Naive W−geodesic (Ng = 4) − Fasano–Franceschini
Sample size (n = m) = W−geodesic (Ng = 3) − W−geodesic (Ng = 5) − AD−geodesic (Ng = 4) − − Upper bound

Effect of neighboring amino-acids on (ϕ, ψ) distribution⁶

^{6.} Javier GONZÁLEZ-DELGADO et al. "Statistical proofs of the interdependence between nearest neighbor effects on polypeptide backbone conformations". In : Journal of Structural Biology 214.4 (2022), p. 107907

Effect of neighboring amino-acids on (ϕ, ψ) distribution⁶

H_0 : Flory's Isolated Pair hypothesis (IPH)⁷

The identities of left and right amino-acids do not have an effect on (ϕ, ψ) distribution

6. Javier GONZÁLEZ-DELGADO et al. "Statistical proofs of the interdependence between nearest neighbor effects on polypeptide backbone conformations". In : Journal of Structural Biology 214.4 (2022), p. 107907

7. Paul. J. FLORY et al. "Statistical mechanics of chain molecules". In: Biopolymers 8.5 (1969), p. 699-700

Javier González-Delgado **Masserstein tests for IDP** 15/11/2022 17/20

Effect of translated codon on (ϕ, ψ) distribution⁸

^{8.} J. GONZÁLEZ-DELGADO et al. Statistical tests to detect differences between codon-specific Ramachandran plots. Submitted. 2022

^{9.} Image : https ://www.nature.com/scitable/topicpage/the-information-in-dna-determines-cellular-function-6523228/

Effect of translated codon on (ϕ, ψ) distribution⁸

Question

Does the identity of the translated codon have an effect on (ϕ, ψ) distribution?

Javier González-Delgado Masserstein tests for IDP 15/11/2022 18 / 20

^{8.} J. GONZÁLEZ-DELGADO et al. Statistical tests to detect differences between codon-specific Ramachandran plots. Submitted. 2022

^{9.} Image : https ://www.nature.com/scitable/topicpage/the-information-in-dna-determines-cellular-function-6523228/

The identity of the translated codon has an effect on (ϕ, ψ) distribution

Comparison of synonymous codon-specific (ϕ, ψ) distributions, considering three different types of **secondary structures** (α -helixes, β -sheets, others).

The identity of the translated codon has an effect on (ϕ, ψ) distribution

Comparison of synonymous codon-specific (ϕ, ψ) distributions, considering three different types of secondary structures (α -helixes, β -sheets, others).

Detect the effect of a sequence mutation on secondary structure (and thus in function)

- Detect the effect of a sequence mutation on secondary structure (and thus in function)
- Assess the quality/convergence of Molecular Dynamics simulations

- \bullet Detect the effect of a sequence mutation on secondary structure (and thus in function)
- Assess the quality/convergence of Molecular Dynamics simulations
- **O** Compare entire sequences and detect outlier amino-acids

- Detect the effect of a sequence mutation on secondary structure (and thus in function)
- Assess the quality/convergence of Molecular Dynamics simulations
- **O** Compare entire sequences and detect outlier amino-acids
- Understand how to design a protein performing a specific function

- Detect the effect of a sequence mutation on secondary structure (and thus in function)
- Assess the quality/convergence of Molecular Dynamics simulations
- **O** Compare entire sequences and detect outlier amino-acids
- Understand how to design a protein performing a specific function

Thank you for your attention !